

## Lecture 1: Wigner's Semicircle Law

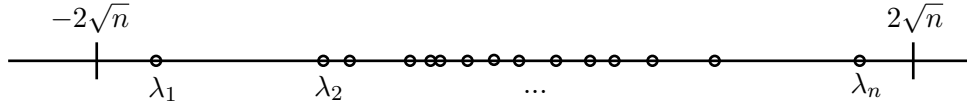
notes by Fan Zhang

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We want to consider **symmetric Wigner matrices**; more precisely, a symmetric Wigner matrix  $M$  has the following properties:

1.  $M$  is symmetric;
2. Each entry of  $M$  is a random variable;
3.  $\{M_{ij}\}_{i \leq j}$  are independent;
4.  $\{M_{ij}\}_{i < j}$  have the same distribution;  $\{M_{ii}\}_{i=1, \dots, n}$  have the same distribution;
5.  $E(M_{ij}) = 0$  for all  $i$  and  $j$ ;
6.  $E(M_{ij}^2) = 1$  for all  $i \neq j$ ;  $E(M_{ii}^2) = 2$  for all  $i = 1, \dots, n$ ;
7.  $M_{ij}$ 's moments are all finite, for all  $i$  and  $j$ .

Then we want to consider the eigenvalues of the matrix  $M$ . Let  $\lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_n$  be the eigenvalues of  $M$ , then they form a point process on  $[-2\sqrt{n}, 2\sqrt{n}]$ . However, there is evidence that the distributions of the eigenvalues are not random; in particular, they are more concentrated around zero.



We will show that when  $n$  is big, the distribution of the eigenvalues approximates a semicircle distribution.

Before we make the theorem precise we need some notations. Let  $M$  be a symmetric Wigner matrix and  $A$  be a set on the real line. Let:

$$\ell^n(A) = E \left( \text{number of eigenvalues of } \frac{1}{\sqrt{n}} M_{n \times n} \text{ in } A \right)$$

Then  $\ell^n$  is obviously a measure. Note that the normalization  $\frac{1}{\sqrt{n}}$  pushes all the eigenvalues to the interval  $[-2, 2]$ . Also, we will define the semicircle distribution by the following density function:

$$f(x) = \frac{1}{2\pi} \sqrt{4 - x^2}; \quad x \in [-2, 2]$$

Note that we renormalized  $M$  by  $\frac{1}{\sqrt{n}}$  because the semicircle distribution on  $[-2, 2]$  has variance 1. We will use  $SC$  to denote the semicircle distribution from now on.

**Theorem 1** (Wigner's Semicircle Law, Average Version).

$$\frac{\ell^n(A)}{n} \xrightarrow{\text{weak}} SC \text{ as } n \rightarrow \infty$$

The proof uses the methods of moments. Let:

$$L^n(A) = \text{number of eigenvalues of } \frac{1}{\sqrt{n}} M_{n \times n} \text{ in } A$$

Where  $A$  is a subset of the real line. Then  $L^n$  is a random measure and it equals to the sum of  $n$  point-masses, i.e.:

$$L^n = \sum_{i=1}^n \delta_{\lambda_i}$$

Furthermore we see that:

$$E(L^n(A)) = \ell^n(A)$$

By construction. We will also use the following notation:

$$\langle f, \mu \rangle = \int f(x) d\mu(x)$$

The idea of the proof is as follows. Let  $k \in \mathbb{N}$  be fixed. Then:

$$\langle x^k, L^n \rangle = \sum_{i=1}^n \lambda_i^k = \text{Tr} \left[ \left( \frac{1}{\sqrt{n}} M_{n \times n} \right)^k \right]$$

We will show that:

$$\langle x^k, \ell^n \rangle = E(\langle x^k, L^n \rangle) = E \left( \text{Tr} \left[ \left( \frac{1}{\sqrt{n}} M_{n \times n} \right)^k \right] \right)$$

Then show that:

$$\frac{1}{n} \langle x^k, \ell^n \rangle \rightarrow \langle x^k, SC \rangle \text{ as } n \rightarrow \infty$$

Once we establish this for any  $k \in \mathbb{N}$ , then we will use the following lemma:

**Lemma 2.** *Let  $\mu_n$  be a sequence of probability measures with bounded support; furthermore,*

$$\langle x^k, \mu_n \rangle \rightarrow \langle x^k, \mu \rangle \text{ as } n \rightarrow \infty$$

*For all  $k \in \mathbb{N}$ . Then*

$$\mu_n \xrightarrow{\text{weak}} \mu$$

Using this lemma we see that

$$\frac{1}{n} \ell^n \xrightarrow{\text{weak}} SC$$

We will now begin the proof of Theorem 1

*Proof.* Consider:

$$\frac{1}{n} \langle x^k, \ell^n \rangle = \frac{1}{n} E \left( \text{Tr} \left[ \left( \frac{1}{\sqrt{n}} M_{n \times n} \right)^k \right] \right) = \frac{1}{n^{1+\frac{k}{2}}} \sum_{i_1, i_2, \dots, i_k=1}^n E(M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1})$$

The key point is to analyze the sum on the right. Note that  $E(M_{ik}) = 0$  for all  $i, j$ . Therefore, in order for the sum on the right to be nonzero, we need to repeat every term at least once; for example, if  $M_{i_1 i_2}$  is different from all of  $M_{i_2 i_3} \dots M_{i_k i_1}$ , then:

$$E(M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}) = E(M_{i_1 i_2}) E(M_{i_2 i_3} \dots M_{i_k i_1}) = 0$$

Now, consider:

$$\begin{aligned} & \frac{1}{n^{1+\frac{k}{2}}} \sum_{i_1, i_2, \dots, i_k=1}^n E(M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}) \\ &= \frac{1}{n^{1+\frac{k}{2}}} \sum_{\text{double path}} E(M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}) + \frac{1}{n^{1+\frac{k}{2}}} \sum_{\text{other paths}} E(M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}) \end{aligned}$$

We will give a combinatorial argument. For paths that do not go over every vertex twice (i.e. the 'other paths'), they can meet at most  $\lfloor \frac{k+1}{2} \rfloor$  vertices. Therefore,

$$\frac{1}{n^{1+\frac{k}{2}}} \sum_{\text{other paths}} E(M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}) \leq \frac{1}{n^{1+\frac{k}{2}}} C_k \frac{n!}{\lfloor \frac{k+1}{2} \rfloor! (n - \lfloor \frac{k+1}{2} \rfloor)!} \leq \frac{1}{n^{1+\frac{k}{2}}} C'_k n^{\lfloor \frac{k+1}{2} \rfloor} \rightarrow 0$$

As  $n \rightarrow \infty$  for some  $C_k, C'_k < \infty$ . For the double path sum, note each path passes through exactly  $k/2$  different vertices twice. There are  $\binom{n}{k/2}$  ways to choose these vertices, and  $m_k$  ways to select the paths once the vertices are fixed, where  $m_k = C_{k/2}$  if  $k$  is even and 0 if  $k$  is odd. Note that  $C_n$  is the Catalan number given by the formula:

$$C_n = \frac{1}{n+1} \binom{2n}{n}$$

(It is an exercise to show that the combinatorial statement about is true.) Therefore,

$$\begin{aligned} & \frac{1}{n^{1+\frac{k}{2}}} \sum_{\text{double path}} E(M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}) = \frac{1}{n^{1+\frac{k}{2}}} m_k n(n-1)(n-2) \dots (n - (k/2 + 1) + 1) \\ & \sim \frac{1}{n^{1+\frac{k}{2}}} m_k n^{k/2+1} = m_k \end{aligned}$$

In summary, we have shown that:

$$\lim_{n \rightarrow \infty} \frac{1}{n} \langle x^k, \ell^n \rangle = m_k$$

Now we want to show that

$$\langle x^k, SC \rangle = m_k$$

To do this, consider a simple random walk on  $\mathbb{Z}$  starting at 0, and let:

$$T = \min \{t : X_t = -1\}$$

Now fix  $z$  with  $|z| < 1$ . Consider:

$$E(z^{T-1}) = g(z) = \sum_{n=0}^{\infty} z^n P(T-1 = n) = \sum_{n=0}^{\infty} \frac{1}{2^{n+1}} z^n m_k$$

If we can show that:

$$g(z) = \sum_{n=0}^{\infty} \frac{1}{2^{k+1}} z^k m_k = \sum_{k=0}^{\infty} \frac{1}{2^{k+1}} z^k S^k = \frac{1}{2} E \left( \frac{1}{1 - \frac{zS}{2}} \right) = \frac{1}{2\pi} \int_{-2}^2 \frac{1}{1 - \frac{zs}{2}} \sqrt{4 - s^2} ds$$

Where  $S \sim SC$ , then  $\langle x^k, SC \rangle = m_k$  for all  $k$  and we complete the proof. To do this, first note that:

$$E(z^T) = zg(z) = \frac{1}{2}z + \frac{1}{2}zE(z^{T_1+T_2})$$

Where  $T_1, T_2, T$  are independent and have the same distributions. Therefore we get:

$$zg(z) = \frac{1}{2}z(1 + (zg(z))^2)$$

Which is a quadratic equation. Solving yields:

$$g(z) = \frac{1 \pm \sqrt{1 - z^2}}{z^2}$$

However, by definition  $g(0) = E(0^{T-1}) = 1$ , so our only solution is:

$$g(z) = \frac{1 - \sqrt{1 - z^2}}{z^2}$$

All there is left to show (exercise) is that:

$$g(z) = \frac{1 - \sqrt{1 - z^2}}{z^2} = \frac{1}{2\pi} \int_{-2}^2 \frac{1}{1 - \frac{zs}{2}} \sqrt{4 - s^2} ds$$

Which will complete the proof. □

**Theorem 3** (Wigner's Semicircle Law).

$$\frac{1}{n}L^n \Rightarrow SC \text{ as } n \rightarrow \infty$$

*Weakly in probability.*

*Proof.* It is sufficient to show that:

$$\frac{1}{n} \langle x^k, L^n \rangle - \frac{1}{n} \langle x^k, \ell^n \rangle \xrightarrow{P} 0$$

For every  $k$ . To do that, we will show that:

$$E \left[ \left( \frac{1}{n} \langle x^k, L^n \rangle \right)^2 \right] - \left( E \left[ \frac{1}{n} \langle x^k, L^n \rangle \right] \right)^2 \rightarrow 0 \text{ as } n \rightarrow \infty$$

Then Chebyshev's Inequality yields the desired result. Note that:

$$\begin{aligned}
& E \left[ \left( \frac{1}{n} \langle x^k, L^n \rangle \right)^2 \right] - \left( E \left[ \frac{1}{n} \langle x^k, L^n \rangle \right] \right)^2 = \frac{1}{n^{2+k}} \left( E \left[ (\text{Tr } M^k)^2 \right] - (E [\text{Tr } M^k])^2 \right) \\
&= \frac{1}{n^{2+k}} \left( E \left[ \left( \sum_{i_1, i_2, \dots, i_k=1}^n (M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}) \right) \left( \sum_{j_1, j_2, \dots, j_k=1}^n (M_{j_1 j_2} M_{j_2 j_3} \dots M_{j_k j_1}) \right) \right] \right. \\
&\quad \left. - \left( \sum_{i_1, i_2, \dots, i_k=1}^n E (M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}) \right) \left( \sum_{j_1, j_2, \dots, j_k=1}^n E (M_{j_1 j_2} M_{j_2 j_3} \dots M_{j_k j_1}) \right) \right) \\
&= \frac{1}{n^{2+k}} \sum_{i \leq j} \text{Cov} \left( \sum_{i_1, i_2, \dots, i_k=1}^n (M_{i_1 i_2} M_{i_2 i_3} \dots M_{i_k i_1}), \sum_{j_1, j_2, \dots, j_k=1}^n (M_{j_1 j_2} M_{j_2 j_3} \dots M_{j_k j_1}) \right) \\
&= \frac{1}{n^{2+k}} C_k n^{k+1} \rightarrow 0 \text{ as } n \rightarrow \infty
\end{aligned}$$

Where the last step requires a combinatorial argument similar to our proof for Theorem 1. □